

Problem 1.

I. Which of the following is a *random* variable?

- a) The population variance
- ✓ b) The sample mean
- c) The variance of the sample mean
- ✓ d) The smallest value in the sample
- ✓ e) The estimated variance of the sample mean
- f) The sample size

II. True or False?

- True g) The center of a 95% confidence interval for the population mean is a random variable
- False h) A 95% confidence interval for μ contains the sample mean with probability 0.95
- False i) A 95% confidence interval contains 95% of the population
- False j) Out of one thousand 95% confidence intervals for μ , 950 will contain μ .

a) The population variance σ^2 is not random ;
it is an unknown constant

b) The sample mean $\bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i$ is random since it depends
on random variables $X_1, \dots, X_n$

c) The variance of the sample mean, $V[\bar{X}_n] = \frac{\sigma^2}{n} \left(1 - \frac{n-1}{N-1}\right)$
is a constant, not a random variable

d) The smallest value in the sample $X_{(1)} = \min \{X_1, \dots, X_n\}$
is random since depends on $X_1, \dots, X_n$

e) The estimated variance of the sample mean

$$s_{\bar{X}_n}^2 = \frac{Nn-n}{Nn-N} \left(1 - \frac{n-1}{N-1}\right) \frac{1}{n} \cdot \left(\frac{1}{n} \sum_{i=1}^n (X_i - \bar{X}_n)^2\right) \text{ is random}$$

f) The sample size n is a constant

g) Confidence interval is random
 $\Rightarrow$ its center is also random

h) i) j) are False, since

a 100(1- α)% confidence interval contains μ with probability (1- α)

a) d) e)

g) is true

h) i) j) are false

Problem 2. Consider a population of size $N=5$, the members of which have values x_1, x_2, x_3, x_4, x_5 .

- 1) If simple random sampling were used, how many samples of size $n=3$ are there?
- 2) Suppose that rather than simple random sampling, the following sampling scheme is used. The possible samples of size $n=3$ are

$$\{x_1, x_2, x_3\}, \{x_2, x_3, x_4\}, \{x_3, x_4, x_5\}, \{x_4, x_5, x_1\}, \{x_5, x_1, x_2\}$$

and the sampling is done in such a way that each of these 5 possible samples is equally likely. Is the sample mean unbiased?

1) How many times we can pick 3 objects out of 5?

We can simply list all of them

$$\left\{ \begin{array}{l} x_1 x_2 x_3 \\ x_1 x_2 x_4 \\ x_1 x_2 x_5 \\ x_1 x_3 x_4 \\ x_1 x_3 x_5 \\ x_1 x_4 x_5 \\ x_2 x_3 x_4 \\ x_2 x_3 x_5 \\ x_2 x_4 x_5 \\ x_3 x_4 x_5 \end{array} \right\}$$

$$\# = \binom{5}{3} = \frac{5!}{3! 2!} = \frac{1 \cdot 2 \cdot 3 \cdot 4 \cdot 5}{1 \cdot 2 \cdot 3 \cdot 1 \cdot 2} = 10$$

or, we can use "n choose k" formula

2) Let us compute the expected value of the sample mean:

$$\begin{aligned} E[\bar{X}_n] &= \frac{x_1 + x_2 + x_3}{3} \cdot \overbrace{P(\{x_1, x_2, x_3\} \text{ is chosen})}^{1/5} + \\ &+ \frac{x_2 + x_3 + x_4}{3} \cdot \overbrace{P(\{x_2, x_3, x_4\} \text{ is chosen})}^{1/5} + \dots + \\ &+ \frac{x_5 + x_1 + x_2}{3} \cdot \overbrace{P(\{x_5, x_1, x_2\} \text{ is chosen})}^{1/5} = \end{aligned}$$

$$= \frac{1}{5} \cdot \frac{1}{3} \left([x_1 + x_2 + x_3] + [x_2 + x_3 + x_4] + \dots + [x_5 + x_1 + x_2] \right) =$$

$$= \frac{1}{5} \cdot \frac{1}{3} \cdot 3 (x_1 + x_2 + x_3 + x_4 + x_5)$$

$$= \frac{x_1 + x_2 + x_3 + x_4 + x_5}{5} = \mu$$

for every i ,
there 3 terms
 x_i in the
sum

$$1) \# = 10$$

$$2) \text{ Yes}$$

Problem 3. Suppose X is a random variable with mean μ and variance σ^2 . Use the delta method to approximate mean and variance of $Y = \log(X)$.

Hint: To approximate $E[Y]$, use the Taylor series expansion of $g(x) = \log(x)$ at $x = \mu$ to the second order; to approximate $\text{Var}[Y]$, use the Taylor series expansion of $g(x) = \log(x)$ at $x = \mu$ to the first order.

1. Taylor expansion of $g(x) = \log(x)$ at μ :

$$g(x) \approx g(\mu) + (x-\mu)g'(\mu) + \frac{1}{2}(x-\mu)^2 g''(\mu)$$

$$g'(x) = \frac{1}{x}$$

$$g''(x) = -\frac{1}{x^2}$$

$$g(x) \approx \log \mu + (x-\mu) \frac{1}{\mu} - \frac{1}{2}(x-\mu)^2 \frac{1}{\mu^2}$$

2. Mean

$$\begin{aligned} E[Y] &= E[g(x)] \approx E\left[\log \mu + (x-\mu) \frac{1}{\mu} - \frac{1}{2\mu^2} (x-\mu)^2\right] \\ &= \log \mu - \frac{1}{2\mu^2} \underbrace{E[(x-\mu)^2]}_{\sigma^2} = \log \mu - \frac{\sigma^2}{2\mu^2} \end{aligned}$$

3. Variance

$$\begin{aligned} V[Y] &= V[g(x)] \approx V\left[\log \mu + (x-\mu) \frac{1}{\mu}\right] \\ &= V\left[\frac{X}{\mu} - 1\right] = \frac{1}{\mu^2} V[X] = \frac{\sigma^2}{\mu^2} \end{aligned}$$

$$\begin{aligned} E[Y] &\approx \log \mu - \frac{\sigma^2}{2\mu^2} \\ V[Y] &\approx \frac{\sigma^2}{\mu^2} \end{aligned}$$

Problem 4. An investigator quantifies her uncertainty about the estimate of a population mean by reporting $\bar{X}_n \pm 2s_{\bar{X}_n}$. What size confidence interval is this?

Hint: It may help to know the following value $\Phi(2) \approx 0.977$, where Φ is of the standard normal CDF.

An approximate $100(1-\alpha)\%$ confidence interval for μ is

$$\bar{X}_n \pm z_{\frac{\alpha}{2}} \sigma_{\bar{X}_n}$$

$S_{\bar{X}_n}^2$ is an unbiased estimate of $\sigma_{\bar{X}_n}^2$.

Thus, in practice, $S_{\bar{X}_n}$ is used instead of $\sigma_{\bar{X}_n}$ (since $\sigma_{\bar{X}_n}$ is typically unknown)

The reported interval is

$$\bar{X}_n \pm 2 S_{\bar{X}_n}$$

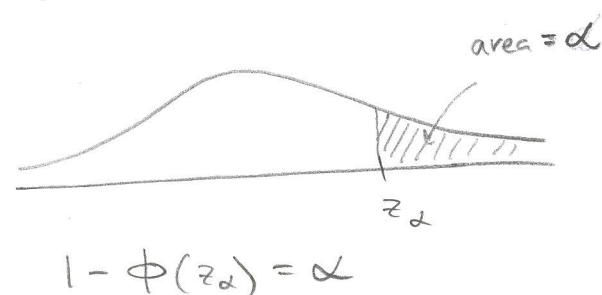
To find α , we have to solve : $z_{\frac{\alpha}{2}} = 2$ for α .

This equation is equivalent to

$$\Phi(z_{\frac{\alpha}{2}}) = \Phi(2)$$

$$1 - \frac{\alpha}{2} = \Phi(2)$$

$$\alpha = 2 \cdot (1 - \Phi(2)) \approx 0.046$$



$\Rightarrow \bar{X}_n \pm 2 S_{\bar{X}_n}$ is an approximate 95.4% confidence interval

size = 95.4%

Problem 5. The following table (Cochran 1977) shows the stratification of all farms in a county by farm size and the mean and standard deviations of the number of acres of corn in each stratum.

Farm Size	N_k	μ_k	σ_k
0-40	394	5.4	8.3
41-80	461	16.3	13.3
81-120	391	24.3	15.1
121-160	334	34.5	19.8
161-200	169	42.1	24.5
201-240	113	50.1	26.0
241 +	148	63.8	35.2

- For a sample size of $n=100$ farms, compute the sample sizes from each stratum for proportional and optimal allocations.
- Calculate the variances of the sample mean for each allocation and compare them to each other

a) For proportional allocation : $\tilde{n}_k = n \cdot w_k = n \cdot \frac{N_k}{N} = n \cdot \frac{N_k}{\sum_{k=1}^7 N_k}$

For optimal allocation : $\hat{n}_k = n \frac{w_k \sigma_k}{\sum_{j=1}^7 w_j \sigma_j}$

Using values from the table we obtain :

k	1	2	3	4	5	6	7
$\tilde{n}_k$	$19.6 \approx \boxed{20}$	$22.9 \approx \boxed{23}$	$19.45 \approx \boxed{19}$	$16.6 \approx \boxed{17}$	$8.4 \approx \boxed{8}$	$5.6 \approx \boxed{6}$	$7.36 \approx \boxed{7}$
$\hat{n}_k$	$9.56 \approx \boxed{10}$	$17.9 \approx \boxed{18}$	$17.26 \approx \boxed{17}$	$19.3 \approx \boxed{19}$	$12.1 \approx \boxed{12}$	$8.6 \approx \boxed{9}$	$15.2 \approx \boxed{15}$

b) $V[\bar{X}_{n,opt}^*] = \frac{1}{n} \left(\sum_{k=1}^7 w_k \sigma_k \right)^2 \approx \underline{2.9}$

$V[\bar{X}_{n,p}^*] = \frac{1}{n} \sum_{k=1}^7 w_k \sigma_k^2 \approx \underline{3.4}$

$V[\bar{X}_{n,p}^*] > V[\bar{X}_{n,opt}^*]$

$\tilde{n}_k = [20, 23, 19, 17, 8, 6, 7]$ $\hat{n}_k = [10, 18, 17, 19, 12, 9, 15]$ $V[\bar{X}_{n,p}^*] = 3.4 \quad V[\bar{X}_{n,opt}^*] = 2.9$
