

Lecture 9-10. Tricks with Random Variables:
The Law of Large Numbers &
The Central Limit Theorem

February 6-8, 2013

Agenda

- Large Sample Theory
- Types of Convergence
 - ▶ Convergence in Probability
 - ▶ Convergence in Distribution
- The Law of Large Numbers
 - ▶ The Monte Carlo Method
- The Central Limit Theorem
 - ▶ Multivariate version
- Summary

Large Sample Theory

The most important aspect of probability theory concerns the **behavior of sequences of random variables**. This part of probability is called **large sample theory** or **limit theory** or **asymptotic theory**. This theory is extremely important for **statistical inference**.

The basic question is this:

What can we say about the limiting behavior of a sequence of random variables?

$$X_1, X_2, X_3 \dots$$

In the statistical context: What happens as we gather more and more data?

In **Calculus**, we say that a **sequence of real numbers** x_1, x_2, \dots converges to a limit x if, for every $\epsilon > 0$, we can find N such that $|x_n - x| < \epsilon$ for all $n > N$.

In **Probability**, **convergence is more subtle**.

Going back to calculus, suppose that $x_n = 1/n$. Then trivially, $\lim_{n \rightarrow \infty} x_n = 0$. Consider a **probabilistic version** of this example: suppose that X_1, X_2, \dots are independent and $X_n \sim \mathcal{N}(0, 1/n)$. Intuitively, X_n is very concentrated around 0 for large n , and we are tempted to say that X_n “converges” to zero. However, $\mathbb{P}(X_n = 0) = 0$ for **all** n !

Types of Convergence

There are two main types of convergence:

convergence in probability and convergence in distribution

Definition

Let X_1, X_2, \dots be a sequence of random variables and let X be another random variable. Let F_n denote the CDF of X_n and let F denote the CDF of X .

- X_n **converges to X in probability**, written $X_n \xrightarrow{\mathbb{P}} X$,
if for every $\epsilon > 0$

$$\lim_{n \rightarrow \infty} \mathbb{P}(|X_n - X| \geq \epsilon) = 0$$

- X_n **converges to X in distribution**, written $X_n \xrightarrow{\mathcal{D}} X$,
if

$$\lim_{n \rightarrow \infty} F_n(x) = F(x)$$

for all x for which F is continuous.

Relationships Between the Types of Convergence

Example: Let $X_n \sim \mathcal{N}(0, 1/n)$. Then

- $X_n \xrightarrow{\mathbb{P}} 0$
- $X_n \xrightarrow{\mathcal{D}} 0$

Question: Is there any relationship between $\xrightarrow{\mathbb{P}}$ and $\xrightarrow{\mathcal{D}}$?

Answer: Yes:

$$X_n \xrightarrow{\mathbb{P}} X \quad \text{implies that} \quad X_n \xrightarrow{\mathcal{D}} X$$

Important Remark: The reverse implication does not hold:
convergence in distribution does not imply convergence in probability.

Example: Let $X \sim \mathcal{N}(0, 1)$ and let $X_n = -X$ for all n . Then

- $X_n \xrightarrow{\mathcal{D}} X$
- $X_n \not\xrightarrow{\mathbb{P}} X$

The Law of Large Numbers

The **law of large numbers** is one of the main achievements in probability. This theorem says that the **mean of a large sample is close to the mean of the distribution**.

The Law of Large Numbers

Let X_1, X_2, \dots be an i.i.d. sample and let $\mu = \mathbb{E}[X_1]$ and $\sigma^2 = \mathbb{V}[X_1] < \infty$. Then

$$\bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i \xrightarrow{\mathbb{P}} \mu$$

Useful Interpretation:

The distribution of \bar{X}_n becomes **more concentrated around μ** as n gets larger.

Example: Let $X_i \sim \text{Bernoulli}(p)$. The **fraction of heads** after n tosses is \bar{X}_n .

According to the **LLN**, $\bar{X}_n \xrightarrow{\mathbb{P}} \mathbb{E}[X_i] = p$. It means that, when n is large, the distribution of \bar{X}_n is tightly concentrated around p .

Q: How large should n be so that $\mathbb{P}(|\bar{X}_n - p| < \epsilon) \geq 1 - \alpha$?

Answer: $n \geq \frac{p(1-p)}{\alpha\epsilon^2}$

The Monte Carlo Method

Suppose we want to calculate

$$I(f) = \int_0^1 f(x) dx$$

where the integration cannot be done by elementary means.

The **Monte Carlo method** works as follows:

- 1 Generate **independent uniform random variables** on $[0,1]$, $X_1, \dots, X_n \sim U[0, 1]$
- 2 Compute $Y_1 = f(X_1), \dots, Y_n = f(X_n)$. Then Y_1, \dots, Y_n are **i.i.d.**
- 3 By the **law of large numbers** \bar{Y}_n should be close to $\mathbb{E}[Y_1]$. Therefore:

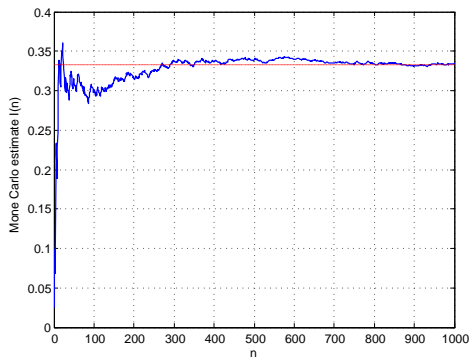
$$\frac{1}{n} \sum_{i=1}^n f(X_i) = \bar{Y}_n \approx \mathbb{E}[Y_1] = \mathbb{E}[f(X_1)] = \int_0^1 f(x) dx$$

Monte Carlo method: Example

Suppose we want to compute the following integral:

$$I = \int_0^1 x^2 dx$$

- From calculus: $I = 1/3$
- Using Monte Carlo method: $I(n) = \frac{1}{n} \sum_{i=1}^n X_i^2$, where $X_i \sim U[0, 1]$



Accuracy of the Monte Carlo method

$$\frac{1}{n} \sum_{i=1}^n f(X_i) \approx \int_0^1 f(x) dx, \quad X_1, \dots, X_n \sim U[0, 1]$$

Question: How large should n be to achieve a desired accuracy?

Answer: Let $f : [0, 1] \rightarrow [0, 1]$. To get $\frac{1}{n} \sum_{i=1}^n f(X_i)$ within ϵ of the true value $I(f)$ with probability at least p , we should choose n so that

$$n \geq \frac{1}{\epsilon^2(1-p)}$$

Thus, the Monte Carlo method tells us how large to take n to get a desired accuracy.

The Central Limit Theorem

Suppose that X_1, \dots, X_n are i.i.d. with mean μ and variance σ^2 . The **central limit theorem** (CLT) says that \bar{X}_n has a distribution which is approximately Normal. This is remarkable since nothing is assumed about the distribution of X_i , except the existence of the mean and variance.

The Central Limit Theorem

Let X_1, \dots, X_n be i.i.d. with mean μ and variance σ^2 . Let $\bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i$. Then

$$Z_n \equiv \frac{\bar{X}_n - \mu}{\sigma/\sqrt{n}} \xrightarrow{\mathcal{D}} Z \sim \mathcal{N}(0, 1)$$

Useful Interpretation:

- Probability statements about \bar{X}_n can be approximated using a Normal distribution.

The Central Limit Theorem

$$Z_n \equiv \frac{\bar{X}_n - \mu}{\sigma/\sqrt{n}} \xrightarrow{\mathcal{D}} Z \sim \mathcal{N}(0, 1)$$

There are several forms of notation to denote the fact that the distribution of Z_n is converging to a Normal. **They all mean the same thing:**

$$Z_n \rightsquigarrow \mathcal{N}(0, 1)$$

$$\bar{X}_n \rightsquigarrow \mathcal{N}\left(\mu, \frac{\sigma^2}{n}\right)$$

$$\bar{X}_n - \mu \rightsquigarrow \mathcal{N}\left(0, \frac{\sigma^2}{n}\right)$$

$$\sqrt{n}(\bar{X}_n - \mu) \rightsquigarrow \mathcal{N}(0, \sigma^2)$$

$$\frac{\bar{X}_n - \mu}{\sigma/\sqrt{n}} \rightsquigarrow \mathcal{N}(0, 1)$$

The Central Limit Theorem: Remarks

- The CLT asserts that the CDF of \bar{X}_n , suitably normalized to have mean 0 and variance 1, converges to the CDF of $\mathcal{N}(0, 1)$.

Q: Is the corresponding result valid at the level of PDFs and PMFs?

Broadly speaking the answer is **yes**, but some condition of smoothness is necessary (generally, $F_n(x) \rightarrow F(x)$ does not imply $F'_n(x) \rightarrow F'(x)$).

- The CLT does not say anything about the **rate of convergence**.
- The CLT tells us that in the long run **we know what the distribution must be**.
 - ▶ Even better: it is **always the same distribution**.
 - ★ Still better: it is one which is **remarkably easy** to deal with, and for which we have a **huge amount of theory**.

Historic Remark:

- For the **special case of Bernoulli variables with $p = 1/2$** , CLT was proved by **de Moivre** around **1733**.
- General values of p were treated later by **Laplace**.
- The first **rigorous proof of CLT** was discovered by **Lyapunov** around **1901**.

The Central Limit Theorem: Example

- Suppose that the number of errors per computer program has a **Poisson distribution** with mean $\lambda = 5$. $f(k|\lambda) = e^{-\lambda} \frac{\lambda^k}{k!}$
- We get $n = 125$ programs; n is **sample size**
- Let X_1, \dots, X_n be the **number of errors in the programs**, $X_i \sim \text{Poisson}(\lambda)$.
- Estimate probability $\mathbb{P}(\bar{X}_n \leq \lambda + \epsilon)$, where $\epsilon = 0.5$.

Answer:

$$\mathbb{P}(\bar{X}_n \leq \lambda + \epsilon) \approx \Phi\left(\epsilon \sqrt{\frac{n}{\lambda}}\right) = \Phi(2.5) \approx 0.994$$

The Central Limit Theorem: Example

- A tourist in Las Vegas was attracted by a certain **gambling game** in which
 - ▶ the customer stakes **1 dollar** on each play
 - ▶ a **win** then pays the customer **2 dollars plus** the return of her **stake**
 - ▶ a **loss** costs her only **her stake**
- The probability of winning at this game is $p = 1/4$.
- The tourist played this game $n = 240$ times.

Assuming that **no near miracles happened**,

- about how much poorer was the tourist upon leaving the casino?

Answer:

$$\mathbb{E}[\text{payoff}] = -\$60$$

- what is the probability that she lost no money?

Answer:

$$\mathbb{P}[\text{payoff} \geq 0] \approx 0$$

The Central Limit Theorem

The **central limit theorem** tells us that

$$Z_n = \frac{\bar{X}_n - \mu}{\sigma/\sqrt{n}} \stackrel{\sim}{\sim} \mathcal{N}(0, 1)$$

However, in applications, we **rarely know** σ . We can **estimate** σ^2 from X_1, \dots, X_n by **sample variance**

$$S_n^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X}_n)^2$$

Question: If we replace σ with S_n is the central limit theorem still true?

Answer: Yes!

Theorem

Assume the same conditions as the CLT. Then,

$$\boxed{\frac{\bar{X}_n - \mu}{S_n/\sqrt{n}} \xrightarrow{\mathcal{D}} Z \sim \mathcal{N}(0, 1)}$$

Multivariate Central Limit Theorem

Let X_1, \dots, X_n be i.i.d. random vectors with mean μ and covariance matrix Σ :

$$X_i = \begin{pmatrix} X_{1i} \\ X_{2i} \\ \vdots \\ X_{ki} \end{pmatrix} \quad \mu = \begin{pmatrix} \mu_1 \\ \mu_2 \\ \vdots \\ \mu_k \end{pmatrix} = \begin{pmatrix} \mathbb{E}[X_{1i}] \\ \mathbb{E}[X_{2i}] \\ \vdots \\ \mathbb{E}[X_{ki}] \end{pmatrix}$$

$$\Sigma = \begin{pmatrix} \mathbb{V}[X_{1i}] & \text{Cov}(X_{1i}, X_{2i}) & \dots & \text{Cov}(X_{1i}, X_{ki}) \\ \text{Cov}(X_{2i}, X_{1i}) & \mathbb{V}[X_{2i}] & \dots & \text{Cov}(X_{2i}, X_{ki}) \\ \vdots & \vdots & \ddots & \vdots \\ \text{Cov}(X_{ki}, X_{1i}) & \dots & \text{Cov}(X_{ki}, X_{k-1i}) & \mathbb{V}[X_{ki}] \end{pmatrix}$$

Let $\bar{X}_n = (\bar{X}_{1n}, \dots, \bar{X}_{kn})^T$. Then

$$\sqrt{n}(\bar{X}_n - \mu) \xrightarrow{\mathcal{D}} \mathcal{N}(0, \Sigma)$$

Summary

- $X_n \xrightarrow{\mathbb{P}} X$: X_n converges to X in probability, if for every $\epsilon > 0$

$$\lim_{n \rightarrow \infty} \mathbb{P}(|X_n - X| \geq \epsilon) = 0$$

- $X_n \xrightarrow{\mathcal{D}} X$: X_n converges to X in distribution, if for all x for which F is continuous

$$\lim_{n \rightarrow \infty} F_n(x) = F(x)$$

- $X_n \xrightarrow{\mathbb{P}} X$ implies that $X_n \xrightarrow{\mathcal{D}} X$
- **The Law of Large Numbers**: Let X_1, X_2, \dots be an i.i.d. sample and let $\mu = \mathbb{E}[X_1]$. Then

$$\bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i \xrightarrow{\mathbb{P}} \mu$$

- **The Central Limit Theorem**: Let X_1, \dots, X_n be i.i.d. with mean μ and variance σ^2 . Then

$$Z_n \equiv \frac{\bar{X}_n - \mu}{\sigma/\sqrt{n}} \xrightarrow{\mathcal{D}} Z \sim \mathcal{N}(0, 1)$$