

Lecture 6: Collusion and Cartels, Part 2

EC 105. Industrial Organization

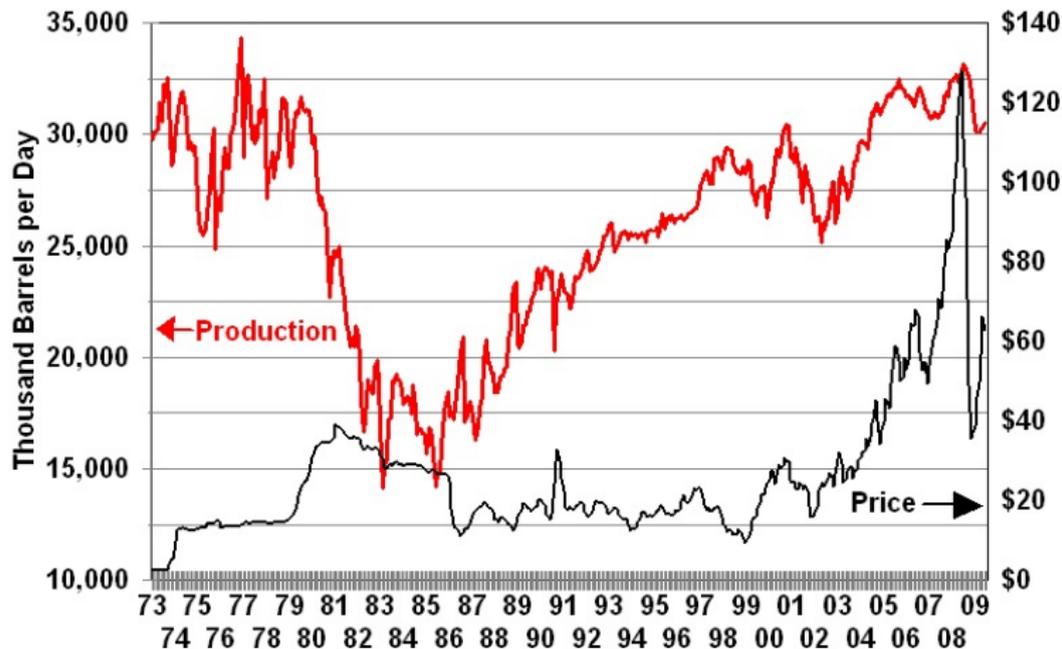
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Outline

Does theory match reality? OPEC

Crude Oil Production (Mbbbl/d)
OPEC Countries



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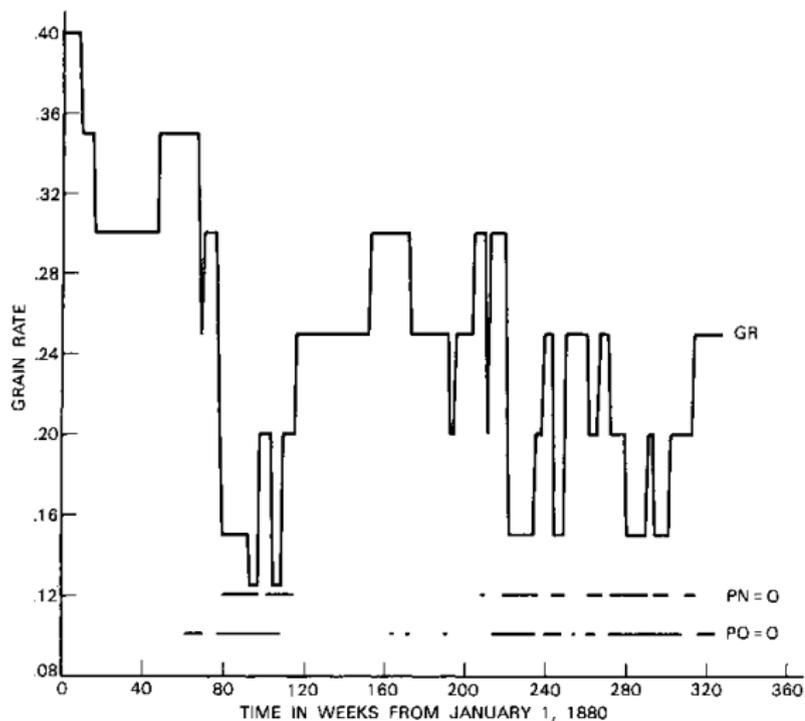
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Does theory match reality? JEC

FIGURE 1

PLOT OF GR, PO, PN AS A FUNCTION OF TIME



Empirical predictions of tacit collusion

- *Constant* production, price
- Does not match empirical and anecdotal evidence from real-world cartels: defection, price-wars, etc.
- Consider two approaches which generates time-varying pricing and production
 - ① Rotemberg-Saloner model: “price war during booms”
 - ② Green-Porter model: secret price cuts
- Important: both models predict time-varying pricing *on the equilibrium path*
 - Changing prices/profitability in any period not due to cheating or punishment.
- Evidence: supermarket pricing
- Case study: Joint Executive Committee (railroad cartel in nineteenth-century US) **[skip]**

Fluctuating Demand: Rotemberg Saloner's (1986) theory of price wars during booms.

- Demand is stochastic.
 1. At each period t , it can be low ($q = D_1(p)$) or high ($q = D_2(p)$) with probability $1/2$ ($D_2(p) > D_1(p)$ for all p). Independent across periods.
 2. At each period firms learn current state before choosing their prices simultaneously: *Bertrand*.
- Look for an optimal stationary symmetric SPNE. A pair of prices $\{p_1, p_2\}$ such that
 1. Firms charge p_s when the state is s ,
 2. Prices $\{p_1, p_2\}$ are sustainable in equilibrium
 3. Expected present discounted profit of each firm along the equilibrium path is Pareto optimal
- Consider infinite stream of payoffs $\pi_0 + \delta\pi_1 + \delta^2\pi_2 + \dots + \delta^n\pi_n + \dots \equiv \Pi (< \infty)$. Then $(1 - \delta)\Pi$ is implied "per-period" payoff. Convenient shorthand in what follows.

Price wars during booms II

- First see whether “fully collusive outcome”, where both firms charge monopoly price p_s^m in each state, is sustainable in eqm.
- Note that payoffs of firm i are, in general:

$$\begin{aligned}\hat{\Pi}^i &= \sum_{t=0}^{\infty} \delta^t \left(\frac{1}{2} \frac{D_1(p_1)}{2} (p_1 - c) + \frac{1}{2} \frac{D_2(p_1)}{2} (p_2 - c) \right) \\ &= \left(\frac{1}{2} \frac{D_1(p_1)}{2} (p_1 - c) + \frac{1}{2} \frac{D_2(p_2)}{2} (p_2 - c) \right) / (1 - \delta)\end{aligned}$$

(Capital Π denotes discounted present value of profit stream.)

- When firms are setting the monopoly prices each period, then the discounted profits (when the the current state is $s \in \{1, 2\}$) is

$$(1 - \delta) \frac{1}{2} \Pi_s^m + \delta \frac{1}{4} (\Pi_1^m + \Pi_2^m)$$

The superscript m denotes monopoly profits.

Price wars during booms III

- It suffices to consider the harshest punishment of switching to competitive price c forever after a deviation (“Bertrand reversion”). If firm i deviates in state s obtains $(1 - \delta)\Pi_s^m + \delta 0$.
- Since $\Pi_1^m < \Pi_2^m$, cheating firms will do so only in state 2; i.e., *incentive constraint* is:

$$(1 - \delta)\Pi_2^m < (1 - \delta)\frac{1}{2}\Pi_2^m + \delta\frac{1}{4}(\Pi_1^m + \Pi_2^m) \quad (*)$$

$$\text{or} \quad \delta > \underline{\delta} \equiv \frac{2\Pi_2^m}{3\Pi_2^m + \Pi_1^m} \in \left[\frac{1}{2}, \frac{2}{3} \right]$$

- Temptation to undercut @ high demand. Compared to stable high demand, face same cheating reward (*-LHS) but lower collusion reward (*-RHS).
- When $\delta \in [1/2, \underline{\delta}]$, full collusion not sustained in high-demand state.
- (Recall: in stationary case, collusion sustainable down to $\delta > \frac{1}{2}$)

Price wars during booms IV

- What is best that firms can do when $\delta \in [1/2, \underline{\delta}]$??
- They can play a **constrained collusion** outcome.
- Choose p_1 and p_2 to: $\max \left(\frac{1}{2} \frac{\Pi_1(p_1)}{2} + \frac{1}{2} \frac{\Pi_2(p_2)}{2} \right)$

- subject to the constraints that for $s = 1, 2$

$$(1 - \delta) \frac{1}{2} \Pi_s(p_s) \leq \delta \frac{1}{4} (\Pi_1(p_1) + \Pi_2(p_2))$$

- Which can be written as:

$$\Pi_1(p_1) \leq \frac{\delta}{2 - 3\delta} \Pi_2(p_2) \quad \text{and} \quad \Pi_2(p_2) \leq \frac{\delta}{2 - 3\delta} \Pi_1(p_1)$$

- As before, the binding constraint is that of state 2. Choosing $p_1 = p_1^m$ increases the objective function and relaxes the constraint for p_2 . Price p_2 is then chosen as high as possible:

$$\Pi_2(p_2) = \frac{\delta}{2 - 3\delta} \Pi_1^m$$

Price wars during booms: Conclusions

- For $\delta \in [1/2, \underline{\delta}]$ some collusion is sustainable.
 1. In the low state of demand, firms charge the monopoly price in that state.
 2. In the high state of demand, firms charge a price below the monopoly price in that state.
- Rotemberg and Saloner interpret this as showing the existence of price war during booms.
 - But note price in high state can be lower or higher than the monopoly price in the low demand state depending on the demand function.
 - This is not a price war in the usual sense, because the price may actually be higher during booms than during busts: we do not obtain from here the implication that oligopoly prices move countercyclically.
 - But *less market power* (lower profitability) during boom periods.

Empirical evidence: Supermarket pricing

- RS predicts that *profitability lower during booms*
- But testing this is challenging: how to separately measure firm profitability and whether economy is in a “boom”?
- Chevalier, Kashyap, Rossi: “Why Don't Prices Rise During Peak Demand?”
 - Consider a number of grocery items.
 - Items have idiosyncratic peak demand periods (tuna/Lent, beer/July4): *not related* to state of economy.
 - Store also has general peak demand periods (Thanksgiving, Christmas)
 - Compare *retail margins* during peak and non-peak demand periods
 - Regression results: look for *negative signs* during peak demand periods.

Products with seasonal demand

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CHEVALIER ET AL.: WHY DON'T PRICES RISE DURING PEAK DEMAND?

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TABLE 4—SEASONAL CHANGES IN RETAIL MARGINS

Panel A: Seasonal Categories							
Variable	Beer	Eating soup	Oatmeal	Cheese	Cooking soup	Snack crackers	Tuna
Linear trend	-0.21 (0.03)	0.04 (0.01)	0.04 (0.01)	-0.009 (0.010)	0.04 (0.01)	0.01 (0.01)	-0.01 (0.01)
Quadratic trend	0.0004 (0.00007)	-0.00009 (0.00002)	-0.00008 (0.00002)	0.00006 (0.00002)	-0.00003 (0.00001)	0.00002 (0.00002)	0.00006 (0.00002)
Cold	-0.01 (0.04)	-0.07 (0.03)	-0.04 (0.02)	0.004 (0.03)	0.02 (0.02)	0.00 (0.03)	-0.02 (0.04)
Hot	-0.03 (0.04)	-0.10 (0.03)	-0.02 (0.02)	-0.08 (0.03)	-0.05 (0.02)	-0.03 (0.03)	-0.07 (0.04)
Lent							-5.03 (1.06)
Easter	0.88 (1.48)	0.34 (1.07)	0.66 (0.65)	-2.57 (1.17)	-1.49 (0.84)	-0.39 (1.20)	-1.82 (1.47)
Memorial Day	-4.36 (1.44)	1.35 (1.11)	1.17 (0.69)	-0.54 (1.22)	0.42 (0.87)	1.59 (1.24)	-1.16 (1.49)
July 4th	-4.08 (1.36)	2.18 (1.18)	0.27 (0.67)	-0.33 (1.29)	0.81 (0.92)	-1.01 (1.31)	1.82 (1.57)
Labor Day	-2.61 (1.33)	1.42 (1.15)	0.19 (0.65)	0.27 (1.25)	0.05 (0.90)	-4.61 (1.28)	-1.50 (1.53)
Thanksgiving	-1.31 (1.54)	1.54 (1.08)	0.01 (0.66)	-5.18 (1.18)	-0.68 (0.84)	-5.04 (1.29)	2.27 (1.36)
Post-Thanksgiving	-3.12 (2.00)	0.87 (1.44)	-1.25 (0.88)	-4.15 (1.59)	0.13 (1.13)	-4.54 (1.72)	0.63 (1.81)
Christmas	-2.66 (1.25)	2.34 (0.90)	-0.42 (0.55)	-3.23 (0.98)	0.51 (0.70)	-8.47 (1.06)	1.00 (1.15)
Constant	28.57 (2.88)	18.42 (0.74)	17.99 (1.00)	34.38 (0.81)	14.21 (0.58)	22.84 (0.83)	25.64 (0.96)
Number of weeks	219	387	304	391	387	385	339

Non-seasonal items

Panel B: Nonseasonal Categories

Variable	Analgesics	Cookies	Crackers	Dish detergent
Linear trend	0.01 (0.00)	0.01 (0.01)	0.02 (0.01)	-0.11 (0.01)
Quadratic trend	0.00002 (0.00001)	0.00001 (0.00002)	0.00000 (0.00002)	0.00028 (0.00002)
Cold	-0.02 (0.02)	0.1 (0.03)	-0.01 (0.03)	0.02 (0.04)
Hot	-0.01 (0.02)	-0.08 (0.03)	-0.03 (0.03)	-0.07 (0.04)
Easter	0.92 (0.73)	-2.71 (1.11)	0.93 (1.10)	1.46 (1.39)
Memorial Day	0.08 (0.75)	1.84 (1.15)	0.16 (1.14)	1.10 (1.44)
July 4th	0.88 (0.80)	1.61 (1.21)	-0.31 (1.21)	0.60 (1.52)
Labor Day	-0.87 (0.78)	1.29 (1.18)	0.58 (1.18)	2.13 (1.49)
Thanksgiving	-0.51 (0.73)	-1.11 (1.19)	0.73 (1.19)	0.53 (1.40)
Post-Thanksgiving	-1.57 (0.98)	-0.53 (1.60)	-0.55 (1.59)	1.99 (1.88)
Christmas	0.50 (0.61)	1.04 (0.95)	0.40 (0.98)	1.16 (1.17)
Constant	25.25 (0.50)	24.14 (0.77)	27.05 (0.77)	27.09 (0.96)
Number of weeks	391	387	385	391

Notes: The dependent variable in each column is the log of the variable-weight retail margin for each category. Units in the table are percentage points. Bold type indicates periods of expected demand peaks. Standard errors are in parentheses.

Secret Price Cuts

- Up to now, we assume that firm's pricing choices are observed by its rival, who can respond by punishing when they observe price cuts. However, (effective) prices may not be observable.
- Must rely on observation of its own realized market share or demand to detect any price undercutting by the rival. But a low market share may be due to the aggressive behavior of one's rival or to a slack in demand.
- Remark: Under uncertainty, mistakes are unavoidable and maximal punishments (eternal reversion to Bertrand behavior) need not be optimal.

Secret Price Cuts

- Framework of our basic repeated game with:
- In each period, there are two possible realizations of demand (states of nature), i.i.d..
 - With probability α , there is no demand for the product sold by the duopolists (the “low-demand” state).
 - With probability $1 - \alpha$, there is a positive demand $D(p)$ (the “high-demand” state).
- A firm that does not sell at some date is unable to observe whether the absence of demand is due to the realization of the low-demand state or to its rival’s lower price.

Secret Price Cuts

- Look for an equilibrium with the following strategies:
 - There is a collusive phase and a punishment phase. The game begins in the collusive phase. Both firms charge p^m until one firm makes a zero profit. (note this is common knowledge).
 - The occurrence of a zero profit triggers a punishment phase. Here both firms charge c for exactly T periods, where T can a priori be finite or infinite.
 - At the end (if any) of the punishment phase, the firms revert to the collusive phase.
- We want to look for a length of the punishment phase such that the expected present value of profits for each firm is maximal subject to the constraint that the associated strategies form a SPNE.

Secret Price Cuts

- Let V^+ denote the present discounted value of a firm's profit from date t on, assuming that at date t the game is in the collusive phase.
- Similarly, let V^- denote the present discounted value of a firm's profit from date t on, assuming that at date t the game is in the punishment phase.
- By the stationarity of the prescribed strategies, V^+ and V^- do not depend on time, and by definition, we have:

$$V^+ = (1 - \alpha) ((1 - \delta)\Pi^m/2 + \delta V^+) + \alpha\delta V^- \quad (1)$$

and

$$V^- = \delta^T V^+ \quad (2)$$

Secret Price Cuts

- Since strategies need to be a SPNE, we need to include *incentive constraints* which rule out profitable deviations in both phases.
- Easy to see there are no profitable 1-shot deviations in punishment phase. Thus, we only consider incentives in the collusive phase:

$$V^+ \geq (1 - \alpha)((1 - \delta)\Pi^m + \delta V^-) + \alpha(\delta V^-) \quad (3)$$

- (3) expresses the trade-off for each firm. If a firm undercuts, it gets $(1 - \delta)\Pi^m > (1 - \delta)\Pi^m/2$. However, undercutting triggers the punishment phase, which yields valuation V^- instead of V^+ .
- To deter undercutting, V^- must be sufficiently lower than V^+ . This means that the punishment must last long enough.
- But because punishments are costly and **occur with positive probability**, T should be chosen as small as possible subj to. Eq. (3)

Secret Price Cuts

- Using (1), we can write (3) as

$$\frac{\delta}{(1-\delta)}(V^+ - V^-) \geq \Pi^m/2 \quad (4)$$

- Also, from (1) and (2) we can get

$$V^+ = \frac{(1-\alpha)(1-\delta)}{(1-(1-\alpha)\delta - \alpha\delta^{T+1})} \frac{\Pi^m}{2} \quad (5)$$

- From (2) we can get $V^+ - V^- = V^+(1 - \delta^T)$, and thus, substituting this and (5) into (4), we can express the incentive constraint as:

$$2(1-\alpha)\delta - \delta^{T+1}(1-2\alpha) \geq 1 \quad (6)$$

Secret Price Cuts

- Note that now we can express the problem as that of maximizing V^+ subject to (6). And furthermore, since V^+ is decreasing in T , we want to find the lowest T such that (6) holds.
- Note that the constraint is not satisfied with $T = 0$, and that therefore, since the LHS of (6) decreases with T if $\alpha \geq 1/2$, that in this case there is no solution (no strategy profile of this sort is a SPNE). Thus we need $\alpha < 1/2$.
- Assuming in fact that $(1 - \alpha)\delta \geq 1/2$, so that the constraint is satisfied for $T \rightarrow \infty$, there exists a (finite) optimal length of punishment T^* . In fact,

$$T^* = \text{int}^+ \left(\frac{\text{Ln} \left(\frac{2(1-\alpha)\delta - 1}{1-2\alpha} \right)}{\text{Ln}(\delta)} - 1 \right)$$

Secret Price Cuts

- This model predicts periodic price wars, contrary to the perfect observation models.
- Price wars are involuntary, in that they are triggered not by a price cut but by an unobservable slump in demand.
 - Price wars purely to provide incentives
 - ... even though firms know that there is no cheating in equilibrium!
(that is, no firm is engaging in secret price-cutting)
- Note also that price wars are triggered by a recession, contrary to the Rotemberg-Saloner model.
 - “price wars during recessions” vs. “price wars during booms”

Secret Price Cuts

- Under imperfect information, the fully collusive outcome cannot be sustained.
 - It could be sustained only if the firms kept on colluding (charging the monopoly price) even when making small profits, because even under collusion small profits can occur as a result of low demand.
 - However, a firm that is confident that its rival will continue cooperating even if its profit is low has every incentive to (secretly) undercut - price undercutting yields a short-term gain and creates no long-run loss.
 - Thus, full collusion is inconsistent with the deterrence of price cuts.

Secret Price Cuts

- Oligopolists are likely to recognize the threat to collusion posed by secrecy, and take steps to eliminate it.
 - Industry trade associations
 1. collect detailed information on the transactions executed by the members.
 2. allows it members to cross-check price quotations.
 3. imposes standardization agreements to discourage price-cutting when products have multiple attributes.
 4. Case study: Joint Executive Committee. Railroad cartel in the late 19th century US.
 - Resale-price maintenance on their retailers, or “most favored nation” clause.
 1. Simplify observation and detection

Porter (1983): Case study of JEC

- Fit data to game theoretic model where behavioral regime – “cooperative” vs. “non-cooperative” – varies over time.
- Reminder: “non-cooperative” phase in repeated games models not due to cheating!
- Measure market power (markup) in both regimes.
- JEC: railroad cartel in 1800s-US.
 - Set price for grain shipments from Midwest to East Coast.
- Data: Table 1
 - Price (GR) and quantity (grain shipments TGR)
 - PO : collusion indicator (could be mismeasured)
 - PN : collusion variable, estimated from model.
 - S_t are *supply-shifters* (dummies DM1, DM2, DM3, DM4 for entry by additional rail companies)
 - $LAKES_t$: dummy when Great Lakes was open to traffic. *Demand shifter* (reduces demand for rail services)

TABLE 1 List of Variables*

<i>GR</i>	grain rate, in dollars per 100 lbs.
<i>TQG</i>	total quantity of grain shipped, in tons.
<i>LAKES</i>	dummy variable; =1 if Great Lakes were open to navigation; =0 otherwise.
<i>PO</i>	cheating dummy variable; =1 if colluding reported by <i>Railway Review</i> ; =0 otherwise.
<i>PN</i>	estimated cheating dummy variable.
<i>DM1</i>	=1 from week 28 in 1880 to week 10 in 1883; =0 otherwise; reflecting entry by the Grand Trunk Railway.
<i>DM2</i>	=1 from week 11 to week 25 in 1883; =0 otherwise; reflecting an addition to New York Central.
<i>DM3</i>	=1 from week 26 in 1883 to week 11 in 1886; =0 otherwise; reflecting entry by the Chicago and Atlantic.
<i>DM4</i>	=1 from week 12 to week 16 in 1886; =0 otherwise; reflecting departure of the Chicago and Atlantic from the JEC.

* The sample is from week 1 in 1880 to week 16 in 1886.

Porter (1983): Model

- N firms (railroads), each producing a homogeneous product (grain shipments). Firm i chooses q_{it} in period t .
- Market demand:

$$\log Q_t = \alpha_0 + \alpha_1 \log p_t + \alpha_2 \text{LAKES}_t + U_{1t}, \quad Q_t = \sum_i q_{it}.$$

- Firm i 's cost fxn: $C_i(q_{it}) = a_i q_{it}^\delta + F_i$
- Firm i 's pricing equation: $p_t(1 + \frac{\theta_{it}}{\alpha_1}) = MC_i(q_{it})$, where:
 - $\theta_{it} =$: 1: (Monopoly/Perfect Collusion); s_{it} (Cournot); 0 (Bertrand)
- After some manipulation, aggregate supply relation is:

$$\log p_t = \log D - (\delta - 1) \log Q_t - \log(1 + \theta_t/\alpha_1)$$

($D = \delta (\sum_i a_i^{1/(1-\delta)})^{1-\delta}$) with empirical version

$$\log p_t = \beta_0 + \beta_1 \log Q_t + \beta_2 S_t + \beta_3 I_t + U_{2t}$$

Porter (1983): Results, Table 3

- Estimate two specifications (differ in whether collusion indicator PO is used)
- GR : price elasticity < 1 in abs. value. Not consistent with optimal monopoly pricing.
- $LAKES_t$ reduces demand;
- DM variables lowered market price
- Coef on PO/PN is > 0 : prices higher when firms are in “cooperative” regime.
 - If we assume that $\theta = 0$ in non-cooperative periods, then this implies $\theta=0.336$ in cooperative periods.
 - Is this too small? (Note $\theta = 1$ w/ perfect cartel)
- Table 4:
 - prices higher and quantity lower in “noncooperative” ($PN = 1$) periods.
 - Cartel earns \$11,000 more in weeks when they are cooperating

TABLE 3 Estimation Results*

Variable	Two Stage Least Squares (Employing PO)		Maximum Likelihood (Yielding PN)**	
	Demand	Supply	Demand	Supply
C	9.169 (.184)	-3.944 (1.760)	9.090 (.149)	-2.416 (.710)
$LAKES$	-.437 (.120)		-.430 (.120)	
GR	-.742 (.121)		-.800 (.091)	
$DM1$		-.201 (.055)		-.165 (.024)
$DM2$		-.172 (.080)		-.209 (.036)
$DM3$		-.322 (.064)		-.284 (.027)
$DM4$		-.208 (.170)		-.298 (.073)
PO/PN		.382 (.059)		.545 (.032)
TQG		.251 (.171)		.090 (.068)
R^2	.312	.320	.307	.863
s	.398	.243	.399	.109

* Monthly dummy variables are employed. To economize on space, their estimated coefficients are not reported. Estimated standard errors are in parentheses.

** PN is the regime classification series (f_1, \dots, f_7). The coefficient attributed to PN is the estimate of β_1 .

TABLE 4 Price, Quantity, and Total Revenue for Different Values of *LAKES* and *PN**

Price	<i>LAKES</i>	
	0	1
<i>PN</i> 0	.1673	.1612
1	.2780	.2679
Quantity	<i>LAKES</i>	
	0	1
<i>PN</i> 0	38680	25904
1	25775	17261
Total Revenue**	<i>LAKES</i>	
	0	1
<i>PN</i> 0	129423	83514
1	143309	92484

* Computed from the reduced form of the maximum likelihood estimates of Table 3, with all other explanatory variables set at their sample means.

** Total Revenue = 20 (Price \times Quantity), to yield dollars per week.

Summary

- Difficulties with taking pure tacit collusion to data: stationary model predicts that prices are unchanging over time
- Rotemberg-Saloner model
 - Perfect information
 - Predictions about varying firm profitability across the business cycle
 - “Price war during booms”
- Green-Porter model
 - Imperfect information; secret price cuts
 - “Price wars during recessions”
- In both cases, price wars arise *not from cheating*, but in order to provide intertemporal incentives to cooperating firms in nonstationary economic environment.