

Non-symmetric Simple Random Walks along Orbits of Ergodic Automorphisms.

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Consider an ergodic measure-preserving automorphism T of probability space (M, \mathcal{M}, μ) . Having a measurable function $p, 0 < p(x) < 1$ a.e., we can construct Markov chain whose phase space is M and a moving point $x \in M$ jumps from x to Tx with probability $p(x)$ and to $T^{-1}x$ with probability $1 - p(x)$. Such Markov chains are called simple random walks along orbits of T . If x_0 is an initial position then the position x_n at any moment of time n can be written as $x_n = T^{b_n}x_0$. For the sequence b_n we have $b_0 = 0$, $b_{n+1} - b_n = \pm 1$, i.e. b_n is a simple random walk on \mathbb{Z}^1 , for which $P\{b_{n+1} = k + 1 | b_n = k\} = p(T^k x)$, $P\{b_{n+1} = k - 1 | b_n = k\} = 1 - p(T^k x)$. The first question which arises in this situation is whether this Markov chain has a stationary measure absolutely continuous wrt to μ . The density r of this measure must satisfy the equation

$$r(x) = r(T^{-1}x)p(T^{-1}x) + r(Tx)(1 - p(Tx)) \quad (1)$$

Assume that $0 < c_1 = \text{const} < p(x) < c_2 = \text{const} < 1$ a.e. and put $q(x) = r(x)p(x)$. Then

$$\frac{q(x)}{p(x)} = q(T^{-1}x) + q(Tx)\frac{1 - p(Tx)}{p(Tx)}$$

or

$$q(x) - q(Tx)\frac{1 - p(Tx)}{p(Tx)} = q(T^{-1}x) - q(x)\frac{1 - p(x)}{p(x)}.$$

This shows that $q(x) - q(Tx)\frac{1 - p(Tx)}{p(Tx)}$ is invariant under T and must be a constant a.e. since T is ergodic:

$$q(x) - q(Tx)\frac{1 - p(Tx)}{p(Tx)} = C$$

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or

$$q(x) = q(Tx) \frac{1 - p(Tx)}{p(Tx)} + C, \quad C = \text{const} \quad (2)$$

Definition 1. Simple random walks is called symmetric if $\int \ln p(x) d\mu(x) = \int \ln(1 - p(x)) d\mu(x)$. Otherwise it is called non-symmetric.

Let us show that in the non-symmetric case (2) and therefore (1) always have a solution. Assume for the definiteness that

$$\int \ln(1 - p(x)) d\mu(x) < \int \ln p(x) d\mu(x). \quad (3)$$

It is easy to see that the series below converges for a.e. x

$$q(x) = C + C \frac{1 - p(Tx)}{p(Tx)} + C \frac{1 - p(Tx)}{p(Tx)} \cdot \frac{1 - p(T^2x)}{p(T^2x)} + \dots \quad (4)$$

Indeed, the product of k terms can be rewritten as

$$\prod_{i=1}^k \frac{(1 - p(T^i x))}{p(T^i x)} = \exp \left[\sum_{i=1}^k \ln \frac{(1 - p(T^i x))}{p(T^i x)} \right].$$

The last sum behaves a.e. as $-ck$ as $k \rightarrow \infty$ in a view of ergodic theorem where $c = \int \ln \frac{p(x)}{1-p(x)} d\mu(x) > 0$. The case $\int \ln(1 - p(x)) d\mu > \int \ln p(x) d\mu(x)$ can be considered in the same way.

If the random walk is symmetric then $C = 0$. Indeed, we must have

$$\int \ln q(x) d\mu(x) = \int \ln \left(q(Tx) \cdot \frac{1 - p(Tx)}{p(Tx)} + C \right) d\mu(x) \quad (5)$$

The right-hand side is a monotone function of C . Since (5) holds for $C = 0$ we get the result. In this case (2) is reduced to

$$\frac{q(x)}{q(Tx)} = \frac{1 - p(Tx)}{p(Tx)}$$

or

$$\ln q(x) - \ln q(Tx) = \ln(1 - p(Tx)) - \ln p(Tx) \quad (6)$$

The last equation has a solution for generic p only if T is a shift on a compact abelian group i.e. it is an automorphism with pure point spectrum [S2] Lect.4. This follows easily from the spectral theory of dynamical systems. This case with $M = Tor^d$ was considered in [S1]. For other transformations having a continuous component in the spectrum the equation (6) has no solution for generic p .

Fix $x \in M$ and consider the probabilities $p_{2n}^{(+)}(x)$ of random walks $\{b(m) \mid 0 \leq m \leq 2n\}$ (see above) for which $b(m) > 0$ for $0 < m < 2n$, $b(2n) = 0$. Following P.Levy we call such walks positive excursions. One can consider also negative excursions and corresponding probabilities $p_{2n}^{(-)}(x)$. It is easy to see that

$$p_2^{(+)}(x) = p(x)(1 - p(Tx)), \quad (7)$$

$$p_{2n}^{(+)}(x) = p(x) \left[\sum_{s \geq 1} \sum_{\substack{n_1 + n_2 + \dots + n_s = n-1 \\ n_i > 0, 1 \leq i \leq s}} p_{2n_1}^{(+)}(Tx) \cdots p_{2n_s}^{(+)}(Tx) \right] (1 - p(Tx)) \quad (8)$$

For the generating $\varphi^{(+)}(x, \Theta) = \sum_{n \geq 1} p_{2n}^{(+)}(x) \Theta^{2n}$, $|\Theta| \leq 1$, we derive from (7) and (8) the equation (see also [S1])

$$\varphi^{(+)}(x, \Theta) = \frac{p(x)(1 - p(Tx))\Theta^2}{1 - \varphi^{(+)}(Tx, \Theta)}. \quad (9)$$

Lemma 1. *If (3) is valid then $\varphi^{(+)}(x, 1) = p(x) \frac{\Lambda(x)}{\Lambda(x)+1}$ where $\Lambda(x) = \sum_{k=1}^{\infty} \prod_{i=1}^k \frac{1-p(T^i x)}{p(T^i x)}$.*

The last series converges a.e.

Proof. For $\varphi^{(+)}(x, 1)$ we have from (9)

$$\varphi^{(+)}(x, 1) = \frac{p(x)(1 - p(Tx))}{1 - \varphi^{(+)}(Tx, 1)}. \quad (10)$$

Then for $\psi^{(+)}(x, 1) = \varphi^{(+)}(x, 1)/p(x) \leq 1$ we have

$$\psi^{(+)}(x, 1) = \frac{\frac{1-p(Tx)}{p(Tx)}}{\frac{1}{p(Tx)} - \psi^{(+)}(Tx, 1)} \quad (11)$$

Since $\psi^{(+)}(y, 1) > 0$ one can write $\psi^{(+)}(T^k x, 1) = \frac{\Lambda(T^k x)}{\Lambda(T^k x) + 1}$. Then (11) gives

$$\frac{\Lambda(x)}{\Lambda(x) + 1} = \frac{\frac{1-p(Tx)}{p(Tx)} (\Lambda(Tx) + 1)}{\frac{1-p(Tx)}{p(Tx)} \Lambda(Tx) + \frac{1}{p(Tx)}}$$

or

$$\Lambda(x) = \frac{1-p(Tx)}{p(Tx)} + \frac{1-p(Tx)}{p(Tx)} \Lambda(Tx) \quad (12)$$

This shows that

$$\Lambda(x) = \sum_{k=1}^{\infty} \prod_{i=1}^k \frac{1-p(T^i x)}{p(T^i x)}$$

The last series converges a.e. in view of (3) and gives the result, QED. Notice that $\Lambda(x) = C^{-1}q(x)$, where $q(x)$ is defined by formula (4).

If we calculate in the same way the distribution of negative Levy excursions we shall easily get for the corresponding generating functions that $\varphi^{(-)}(x, 1) = 1 - p(x)$.

Thus in the non-symmetric case (3) the probability of the negative excursion is $1-p(x)$, the probability of the positive excursion is $p(x) \frac{\Lambda(x)}{\Lambda(x)+1}$ and the probability that the moving point jumps from x to Tx and does not return back to x is $p(x) \frac{1}{\Lambda(x)+1}$. Let us calculate the expectation of the length of the positive excursion $m_1^{(+)}(x) = \sum_{n \geq 1} 2np_{2n}^{(+)}(x)$. We have by differentiating (9)

$$m_1^{(+)}(x) = \frac{\partial \varphi^{(+)}(x, 1)}{\partial \Theta} = \frac{2p(x)(1-p(Tx))}{1-\varphi^{(+)}(Tx, 1)} + \frac{p(x)(1-p(Tx))}{(1-\varphi^{(+)}(Tx, 1))^2} m_1^{(+)}(Tx) \quad (13)$$

The ratio $\frac{p(x)(1-p(Tx))}{(1-\varphi^{(+)}(Tx, 1))^2}$ equals

$$\begin{aligned} \frac{p(x)(1-p(Tx))(\Lambda(Tx) + 1)^2}{((1-p(Tx))\Lambda(Tx) + 1)^2} &= \frac{p(x)}{p(Tx)} \cdot \frac{1-p(Tx)}{p(Tx)} \cdot \frac{(\Lambda(Tx) + 1)^2}{\left(\frac{1-p(Tx)}{p(Tx)} \cdot \Lambda(Tx) + \frac{1-p(Tx)}{p(Tx)} + 1\right)^2} = \\ &= \frac{p(x)}{p(Tx)} \cdot \frac{(\Lambda(Tx) + 1)^2}{(\Lambda(x) + 1)^2} \cdot \frac{1-p(Tx)}{p(Tx)}. \end{aligned}$$

The last step was based on (12). If we put $h(x) = \frac{2p(x)(1-p(Tx))}{1-\varphi^{(+)}(Tx,1)}$ then we can write $m_1^{(+)}(x)$ in the form

$$m_1^{(+)}(x) = h(x) + \frac{p(x)}{(\Lambda(x) + 1)^2} \cdot \sum_{k=1}^{\infty} h(T^k x) \prod_{i=1}^k \frac{1-p(T^i x)}{p(T^i x)} \cdot \frac{(\Lambda(T^k x) + 1)^2}{p(T^{k+1} x)} \quad (14)$$

In view of our assumption concerning p the function h is bounded from above and below. We need

Assumption 1. *The function $\Lambda \in L^1(M, \mathcal{M}, \mu)$.*

Under this assumption the density $r(x) \in L^1(M, \mathcal{M}, \mu)$ (see (1) and (4)).

In a similar way one can write down the expressions for the derivatives $m_r^{(+)}(x) = \frac{d^r \varphi^{(+)}(x,1)}{d\Theta^r}$ which are finite a.e. if the Assumption 1 holds.

The same analysis can be applied to $\varphi^{(-)}(x, 1)$ and its derivatives. One can show that if (3) holds then the probabilities $p_{2n}^{(-)}(x)$ decay exponentially. The sum $\varphi(x, \Theta) = \varphi^{(+)}(x, \Theta) + \varphi^{(-)}(x, \Theta)$ is generating function for excursions, both negative and positive.

Consider the probability distribution $P_x^{(n)}$ concentrated on the set $\bigcup_{|k| \leq n} T^k x$ where

$P_x^{(n)}(k)$ is the probability of random walks starting from $x = x_0$ for which $x_n = T^k x_0$ or $b(n) = k$. For $k > 0$ such random walks have the following structure. First it makes some numbers ν_0 of positive or negative excursions coming back to x_0 . Then it jumps to Tx_0 and makes ν_1 positive excursions. After this it jumps to $T^2 x_1$, makes there ν_2 positive excursions and so on. We denote $\nu_0, \nu_1, \dots, \nu_k$ the numbers of these excursions and $\xi_{i1}, \xi_{i2}, \dots, \xi_{i\nu_i}$ are their lengths, $i = 0, 1, \dots, k$ $\zeta_i = \xi_{i1} + \dots + \xi_{i\nu_i}$. Then $P\{\nu_i = r\} = \left(\frac{p(T^i x) \Lambda(T^i x)}{\Lambda(T^i x) + 1} \right)^r$, $r = 1, 2, \dots$ and we can write

$$P_x^{(n)}(k) = \frac{1}{p(T^k x_0)} \cdot \prod_{i=1}^k p(T^i x) \cdot \frac{1}{1 - \frac{p(T^i x) \Lambda(T^i x)}{\Lambda(T^i x) + 1}} \cdot \tilde{\mathcal{P}}\{\zeta_0 + \zeta_1 + \dots + \zeta_k = n - k\} \quad (15)$$

The factor $\frac{1}{1 - \frac{p(T^i x) \Lambda(T^i x)}{\Lambda(T^i x) + 1}} = \sum_{r=0}^{\infty} \left(\frac{p(T^i x) \Lambda(T^i x)}{\Lambda(T^i x) + 1} \right)^r$ is the sum over the values of ν_i of probabilities to have r positive excursions. From (12)

$$\frac{p(T^i x)}{1 - \frac{p(T^i x) \Lambda(T^i x)}{\Lambda(T^i x) + 1}} = \frac{\Lambda(T^i x) + 1}{\frac{(1-p(T^i x))}{p(T^i x)} (\Lambda(T^i x) + 1)} = \frac{\Lambda(T^i x) + 1}{\Lambda(T^{i-1} x) + 1}$$

Therefore

$$P_x^{(n)}(k) = \frac{1}{p(T^k x)} \frac{\Lambda(T^k x) + 1}{\Lambda(x) + 1} \tilde{\mathcal{P}}\{\zeta_0 + \zeta_1 + \cdots + \zeta_k = n - k\} = \frac{r(T^k x)}{\Lambda(x) + 1} \tilde{\mathcal{P}}\{\zeta_0 + \zeta_1 + \cdots + \zeta_k = n - k\} \quad (16)$$

Here $\tilde{\mathcal{P}}$ is the probability distribution corresponding to the distribution of the sum of $k+1$ independent random variables ζ_i for which $\tilde{\mathcal{P}}\{\zeta_i = 2m\} = \left(1 - \frac{p(T^i x)\Lambda(T^i x)}{\Lambda(T^i x)+1}\right) \mathcal{P}\{\zeta_i = 2m\}$.

Now we can formulate the main result of this paper.

Main theorem. *For a non-symmetric random walks along orbits of T for which inequality (3), Assumption 1 hold and $m_1^{(+)}(x), m_2^{(+)}(x), m_3^{(+)}(x) \in L^1(M, \mathcal{M}, \mu)$ the distribution $P_x^{(n)}$ converges to $r(z)d\mu(z)$ in the following sense: for any bounded measurable φ*

$$\lim_{n \rightarrow \infty} \int \varphi(z) dP_x^{(n)}(z) = \lim_{n \rightarrow \infty} \sum_{k \geq 0} P_x^{(n)}(k) \varphi(T^k x) = \int_M \varphi(z) r(z) d\mu(z)$$

for μ - a.e. x .

Proof. Using (16) we can write

$$\sum_k P_x^{(n)}(k) \varphi(T^k x) = \frac{1}{\Lambda(x) + 1} \sum_k r(T^k x) \varphi(T^k x) \tilde{\mathcal{P}}\{\zeta_0 + \zeta_1 + \cdots + \zeta_k = n - k\}$$

In view of our assumptions the sequence of random variables $\zeta_0, \zeta_1, \dots, \zeta_k$ satisfies the local central limit theorem of probability theory. Therefore if $a(T^i x)$ is the expectation of ζ_i (with respect to $\tilde{\mathcal{P}}$), $\sigma(T^i x)$ is its variance we can write

$$\tilde{\mathcal{P}}\{\zeta_0 + \zeta_1 + \cdots + \zeta_k = n - k\} \sim \frac{1}{\sqrt{2\pi\sigma k}} \cdot \exp \left\{ -\frac{1}{2} \frac{(n - k - \sum_{i=1}^k a_i(T^i x))^2}{\sum_{i=1}^k \sigma(T^i x)} \right\},$$

$\sigma = \frac{1}{k} \sum_{i=1}^k \sigma(T^i x)$. The last formula shows that $\tilde{\mathcal{P}}$ is a slowly varying function of $n - k$. Therefore the average over k of $r(T^k x) \varphi(T^k x)$ with respect to this probability distribution is equivalent to the usual average with respect to the uniform distribution which gives in the limit $\int r(z) \varphi(z) d\mu(z)$, QED.

Theorem (Corollary from the Main Theorem) Let the probability space (M, \mathcal{M}, μ) be the d -dimensional torus Tor^d with the Lebesgue measure $\mu = Leb_d$ and let $T : Tor^d \rightarrow Tor^d$ be a rotation by an irrational angle ω , i.e. $T : x \mapsto x + \omega$. If $p : M \rightarrow (0, 1)$ be a continuous function such that $\int \ln(1 - p(x)) d\mu(x) \neq \int \ln p(x) d\mu(x)$, then the distribution $P_x^{(n)}$ converges to $r(z)d\mu(z)$ in the following sense

$$\lim_{n \rightarrow \infty} \int \phi(z) dP_x^{(n)}(z) = \lim_{n \rightarrow \infty} \sum_{k=1}^n P_x^{(n)}(k) \phi(T^k x) = \int \phi(z) r(z) d\mu(x)$$

for all $x \in M$.

This theorem was proved in [S1] for any diophantine ω .

Proof Assume for definiteness that $\int \ln(1 - p(x)) d\mu(x) < \int \ln p(x) d\mu(x)$.

Proposition $\Lambda(x) = \sum_{k=1}^{\infty} \prod_{i=1}^k \frac{1-p(T^i x)}{p(T^i x)}$ is a continuous function.

Proof of the Proposition. Put $\eta(x) = \ln(1 - p(x))/p(x)$ and $-C = \int \eta(x) d\mu(x) < 0$. Rewrite $\Lambda(x) = \sum_{n=1}^{\infty} \exp(\sum_{k=1}^n \eta(T^k x))$. Show that for any $\epsilon > 0$ there exists a constant $n_0 = n_0(\epsilon)$ such that for any $x \in M$ and $n > n_0$ the following holds:

$$\sum_{k=1}^n \eta(T^k x) < -(C - \epsilon)n < -Cn/2. \quad (17)$$

Since, $\eta(x)$ is a continuous function for any $\epsilon > 0$ there exists $\delta = \delta(\epsilon)$ such that $|\eta(x) - \eta(y)| < \epsilon/2$ for any $|x - y| < \delta$. Fix a δ -grid $\{x_j\}_{j=1}^N$ on $M = Tor^d$. By ergodic theorem there exists a universal constant n_0 such that for any $j = 1, \dots, N$ and any $n > n_0$

$$\left| \frac{1}{n} \sum_{i=1}^n \eta(T^i x_j) - \int \eta(y) d\mu(y) \right| < \epsilon/2.$$

Then for any point $x \in M$ there exists a point of the grid x_j such that $|x - x_j| \leq \delta$. Thus,

$$\left| \frac{1}{n} \sum_{i=1}^n \eta(T^i x) - \int \eta(y) d\mu(y) \right| < \left| \frac{1}{n} \sum_{i=1}^n \eta(T^i x) - \int \eta(y) d\mu(y) \right| + \left| \frac{1}{n} \sum_{i=1}^n (\eta(T^i x) - \eta(T^i x_j)) \right| < \epsilon.$$

Formula (17) implies continuity of $\Lambda(x)$, because $a_n(x) = \exp(\sum_{k=1}^n \eta(T^k x))$ exponentially small and the sum over $n > n_0$ is uniformly small. This completes the proof of the Proposition. QED.

Let us rewrite

$$m_1^{(+)}(x) = \frac{\partial \phi^+(x, \theta)}{\partial \theta} \Big|_{\theta=1}$$

in the form

$$m_1^{(+)}(x) = h(x) + h(x) \sum_{k=1}^{\infty} \prod_{j=0}^k \frac{\Lambda(T^{j+1}x)}{\Lambda(T^jx) + 1}.$$

$\Lambda(x)$ is continuous and, therefore, bounded. This implies that $m^{(+)}(x)$ is a bounded function of x . One can check that

$$m_2^{(+)}(x) = \left. \frac{\partial^2 \phi^+(x, \theta)}{\partial \theta^2} \right|_{\theta=1}$$

can be presented in the following form: Put $\Phi(x) = 1 + \sum_{k=1}^{\infty} \prod_{j=0}^k \frac{\Lambda(T^{j+1}x)}{\Lambda(T^jx) + 1}$. Then

$$m_2^{(+)}(x) = \left(h(x) + 2h(x)h(Tx) \frac{\Lambda(Tx) + 1}{p(Tx)(\Lambda(x) + 1)} \Phi(Tx) \right) \Phi(x).$$

It is easy to see that $\Phi(x)$ is continuous and bounded. Therefore, $m_2^{(+)}(x)$ is bounded. In a similar fashion one can check the third moment. Thus, we can apply the Main theorem. This completes the proof of the theorem. QED.

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Reference.

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